

# Statistics 301: Probability and Statistics

## Continuous Distributions

### Module 5

Updated 2019

## Review of discrete random distributions

From Module 4, the distributions were discrete. A functions associated with a discrete random variable is usually called a *probability mass function (pmf)*. The name pmf is suggested by a model used in physics for a system of “point masses”. The pmf describes how the total probability mass of 1 is distributed at various points along the axis of possible values of the random variable.

## Continuous distributions

### Continuous Random Variable

- A random variable  $X$  is called **continuous** if it can take any value within a finite or infinite interval of the real number line  $(-\infty, \infty)$
- Some examples would be measurements of length, strength, lifetime, pH, etc.

## Density Function of a Continuous Random Variable

A continuous random variable  $X$  cannot have a PMF (probability mass function). The reason for this is that:

$$P(X = x) = 0$$

That is, there is no area under a curve at a single point

## Probability Density Function (pdf)

Probability Density Function (pdf):

The probability density function (pdf) of a continuous random variable  $X$  is a nonnegative function  $f_X$  with the property that  $P(a < X < b)$  equals the area under it and above the interval  $[a, b]$ . Thus,

$$P(a < X < b) = \text{area under } f_X$$

$$a \leq X \leq b$$

## Generic pdf

The area under a curve is found by integration, such as:

$$P(a < X < b) = \int_a^b f(x) dx$$

Keep in mind that the above is equal to the following as you cannot find probabilities of a single point:

$$P(a \leq X \leq b) = \int_a^b f(x) dx$$

## Rules of Probability

1.  $0 \leq p(x) \leq 1$
2.  $\int_{-\infty}^{\infty} f(x) dx = 1$
3. Complement Rule
4. Addition Rule (for disjoint and non-disjoint events)
5. Multiplication Rule (for independent events)
6. Conditional Probability Rule

All the previously learned rules apply to continuous distributions.

## Cumulative Distribution Function

The cumulative distribution function, or **CDF**, of a random variable  $X$  gives the probability of events of the form  $[X \leq x]$ , for all numbers  $x$

*Notation for the cumulative distribution function is:*

CDF or  $F_X(x) = P(X \leq x)$

$$F_X(x) = \int_{-\infty}^x f(y) dy$$

## $EX, VX, SDX$ of Generic pdf

- Expected Value (mean), variance, and standard deviation

$$EX = \int_{-\infty}^{\infty} x f(x) dx$$

$$VX = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$

$$= E(X^2) - (EX)^2 \text{ where } E(X^2) = \int_{-\infty}^{\infty} x^2 f(x) dx$$

$$SDX = \sqrt{VX}$$

## Generic Example I

Let  $X$  denote the resistance of a randomly chosen resistor and suppose its pdf is given as:

$$f(x) = \begin{cases} kx & 8 \leq x \leq 10 \\ 0 & \text{otherwise} \end{cases}$$

- (1) Calculate  $k$
- (2) CDF of  $X$
- (3) Calculate  $P(X < 9)$
- (4) Use CDF to calculate  $P(8.6 \leq X \leq 9.8)$  and  $P(X \leq 9.8 | X \geq 8.6)$
- (5) Calculate  $EX, VX, SDX$

## Generic Example II

$$\begin{aligned} & \int_8^{10} kx dx \\ &= \frac{1}{\frac{1}{2}(10^2 - 8^2)} \Rightarrow k = \frac{1}{18} \end{aligned}$$

$$f(x) = \begin{cases} \frac{1}{8}x & 8 \leq x \leq 10 \\ 0 & \text{otherwise} \end{cases}$$

**The CDF:**

$$F_X(x) = \int_8^x \frac{1}{18}y \, dy \Rightarrow \frac{1}{18}\left(\frac{y^2}{2}\right)\Big|_8^x \\ \Rightarrow \frac{1}{36}(x^2 - 64) = \frac{x^2-64}{36}$$

### Generic Example III

- $P(X < 9)$

$$= \int_8^9 f(x)dx = \int_8^9 \frac{x}{8} dx = F(9) = \frac{9^2-64}{36} = 0.472222$$

### Generic Example IV

- $P(8.6 \leq X \leq 9.8)$

$$= \int_{8.6}^{9.8} f(x)dx = F(9.8) - F(8.6) = \frac{9.8^2-64}{36} - \frac{8.6^2-64}{36} = \\ = 0.613333$$

- For  $P(X \leq 9.8 | X \geq 8.6)$ , formula for conditional probability:  $P(A|B) = \frac{P(A \cap B)}{P(B)}$  so

$$P(X \leq 9.8 | X \geq 8.6) = \frac{P(X \leq 9.8 \cap X \geq 8.6)}{P(X \geq 8.6)} \\ = \frac{P(8.6 \leq X \leq 9.8)}{P(X \geq 8.6)} \\ = \frac{F(9.8) - F(8.6)}{F(10) - F(8.6)} = 0.8479$$

### Generic Example V

- Find  $EX$ ,  $VX$ ,  $SDX$

$$EX = \int_8^{10} x \left(\frac{1}{18}x\right) dx = \frac{1}{18} \int_8^{10} x^2 dx \\ = \frac{1}{18} \left(\frac{x^3}{3}\right)\Big|_8^{10} = \frac{10^3-8^3}{54} = 9.037037$$

$$E(X^2) = \int_8^{10} x^2 \left(\frac{1}{18}x\right) dx = 82$$

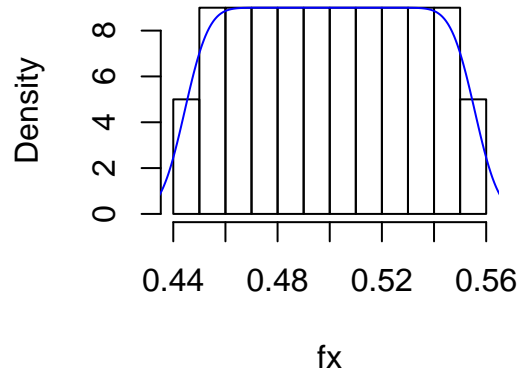
$$VX = E(X^2) - (EX)^2 = 0.331962$$

$$SDX = \sqrt{VX} = 0.576161$$

### Graph of Generic Example pdf

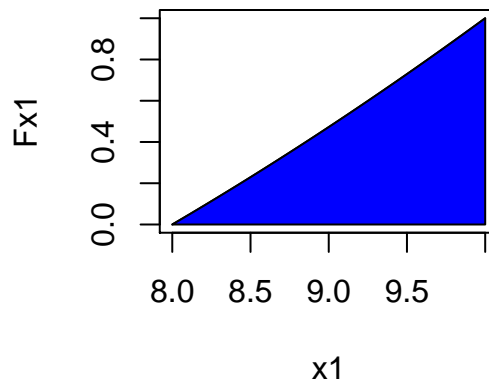
Since there are an infinite number of values within the interval from 8 to 10, I just chose 1000 values (it could have been smaller but I just chose 1000)

### Histogram of fx



### Graph of Generic Example CDF

#### Plot of CDF



### Uniform Distribution

In the context of probability distributions, a uniform distribution refers to a probability distribution for which all of the values that a random variable can take on occur with equal probability. This probability distribution is defined as follows.

A random variable  $X$  is said to be uniform if all values of  $X$  are equally likely

$$X \sim U(A, B)$$

$$P(a < X < b) = \int_a^b \frac{1}{B - A} dx \quad \text{for } A < X < B$$

### Uniform $EX$ , $VX$ , $SDX$

$$EX = \frac{B + A}{2}$$

$$VX = \frac{(B - A)^2}{12}$$

$$SDX = \sqrt{VX}$$

## Uniform Example I

Say that  $Y$  has a uniform distribution on the interval  $[2, 5]$ . Find the following:

- $f(y)$
- $F_Y(y)$
- $P(2 < Y < 3)$
- $EX, VX, SDX$

## Uniform Example II

$$f(y) = \begin{cases} \frac{1}{5-2} & 2 \leq y \leq 5 \\ 0 & \text{otherwise} \end{cases}$$

$$F_Y(y) = \int_2^y \frac{1}{3} dx = \frac{1}{3}(x)_2^y =$$

$$\frac{y-2}{3}$$

$$P(2 < Y < 3)$$

$$= F(3) - F(2) = 0.333333$$

## Uniform Example III

$$EX = \frac{B + A}{2} = \frac{5 + 2}{2} = 3.5$$

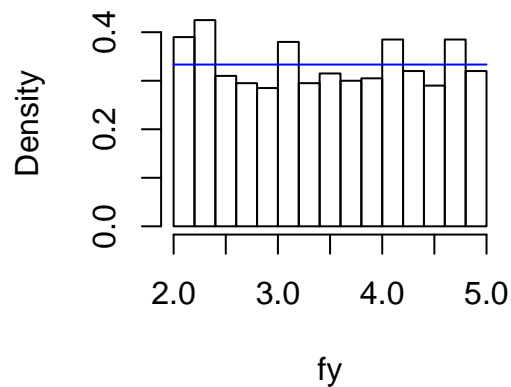
$$VX = \frac{(B - A)^2}{12} = \frac{(5 - 2)^2}{12} = 0.75$$

$$SDX = \sqrt{VX} = \sqrt{0.75} = 0.866025$$

## Graph of Uniform pdf

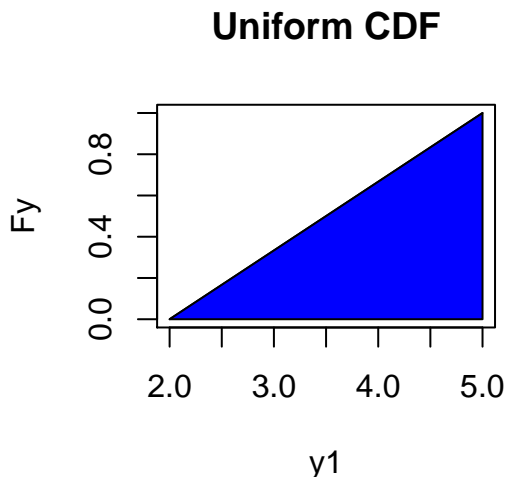
```
y1=seq(from=2,to=5,length.out=1000); fy=runif(length(y1),min=2,max=5)
hist(fy,prob=T); curve(dunif(x,min=2,max=5),col='blue',add=T)
```

**Histogram of fy**



## Graph of Uniform CDF

```
y1=seq(from=2,to=5,length.out=1000); F=antiD((1/3)-y); Fy=F(y1)-F(2)
plot(y1,Fy,type='l',main="Uniform CDF")
polygon(c(y1,y1[length(y1)]), c(Fy,Fy[1]),col='blue')
```



## The Exponential Distribution

$X$  is said to have an **exponential distribution** with parameter  $\lambda$  and pdf:

$$f(x; \lambda) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

and will always have the CDF (through integration by parts):

$$F_X(x) = 1 - e^{-\lambda x}$$

$$EX = \frac{1}{\lambda}$$

$$VX = \frac{1}{\lambda^2}$$

$$SDX = \sqrt{VX}$$

## Exponential Example I

Suppose that the useful time (in years) of a PC is exponentially distributed with parameter  $\lambda = 0.25$ . A student entering a four-year undergraduate program inherits a two-year old PC from his sister who just graduated. Find the probability that the useful lifetime of the PC will last at least until he graduates (assume within 4 years). Let  $X$  denote the useful lifetime of the PC.

$$f(x; \lambda) = f(x; 0.25) = \frac{1}{4}e^{-\frac{1}{4}x}$$

$$P(X > 4 + 2 | X > 2) = \frac{P(X > 4 + 2 \cap X > 2)}{P(X > 2)}$$

$$= \frac{P(X > 6)}{P(X > 2)} = \frac{e^{-.25*6}}{e^{-.25*2}} = 0.367879$$

## Exponential Example II

About how long would you expect the PC to last, on average? (this is the question to find the mean). Find  $EX$ ,  $VX$ ,  $SDX$

$$EX = \frac{1}{\lambda} = \frac{1}{0.25} = 4$$

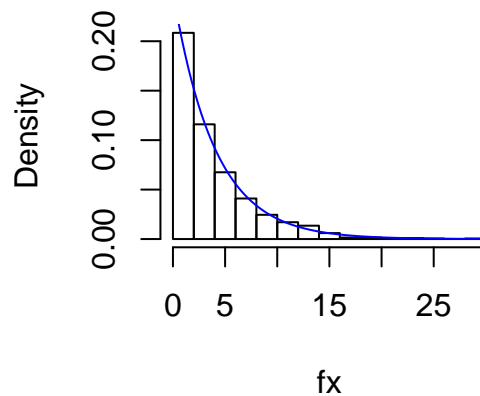
$$VX = \frac{1}{\lambda^2} = \frac{1}{(0.25)^2} = 16$$

$$SDX = \sqrt{VX} = 4$$

## Graph of Exponential pdf

```
fx=rexp(1000,rate=.25)
hist(fx,prob=T,main="Exponential pdf")
curve(dexp(x,rate=.25),col='blue',add=T)
```

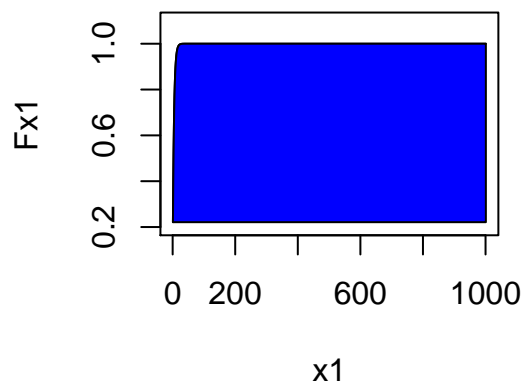
### Exponential pdf



## Graph of Exponential CDF

```
x1=seq(0:1000); F=antiD((1/4)*exp(-.25*x)~x); Fx1=F(x1)-F(0)
plot(x1,Fx1,type='l',main="Exponential CDF",ylim=c(0.2,1.1))
polygon(c(x1,x1[length(x1)]),c(Fx1, Fx1[1]),col='blue')
```

### Exponential CDF



## The Normal Distribution

The normal distribution is one of the most important and widely used. Many populations have distributions that can be fit very closely by an appropriate normal curve.

A continuous rv  $X$  is said to have a **normal distribution** with parameters  $\mu$  and  $\sigma^2$  where  $-\infty < \mu < \infty$  and  $\sigma^2 > 0$ , if the pdf of  $X$  is:

$$f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$$

With  $EX = \mu$ ,  $VX = \sigma^2$  and  $SDX = \sigma$

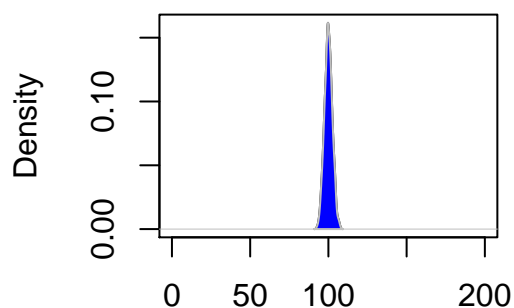
## The Normal Density Curve

The following graphs will illustrate differences in the exact shape of the normal distribution, depending on the standard deviation (or variance).  $\mu$  will be located in the center of the distribution (because of its symmetry) and  $\sigma$  will horizontally extend from  $\mu$  to the first inflection point on the curve. Large values of  $\sigma$  yield graphs that are quite spread out about  $\mu$  (and a value of  $X$  far from  $\mu$  may be well observed), whereas small values of  $\sigma$  yield graphs with a high peak above  $\mu$  and most of the area under the graph quite close to  $\mu$  (implying that a value of  $X$  far from  $\mu$  is quite unlikely).

### Normal Distribution mean=100, sd=2.5

```
d1=density(rnorm(1000,100,2.5))
plot(d1,xlim=c(0,200))
polygon(d1, col="blue",border="grey")
```

**density.default(x = rnorm(1000, 100, 2.5))**



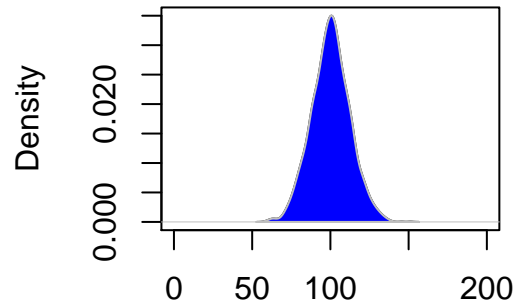
N = 1000 Bandwidth = 0.5668

### Normal Distribution mean=100, sd=12

```
d2=density(rnorm(1000,100,12))
plot(d2,xlim=c(0,200))
polygon(d2, col="blue",border="grey")
```



**ensity.default(x = rnorm(1000, 10**

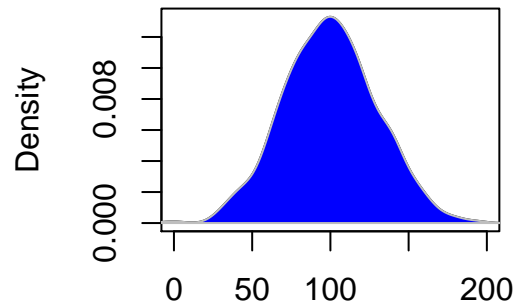


N = 1000 Bandwidth = 2.705

Normal Distribution mean=100, sd=29

```
d3=density(rnorm(1000,100,29))
plot(d3,xlim=c(0,200))
polygon(d3, col="blue",border="grey")
```

**ensity.default(x = rnorm(1000, 10**



N = 1000 Bandwidth = 6.691

**Normal pdf**

When  $X$  is a normal rv with mean  $\mu$  and variance  $\sigma^2$ ,

$$X \sim N(\mu, \sigma)$$

To compute the probabilities, this requires techniques beyond the usual methods; for  $\mu = 0$  and  $\sigma = 1$ , tables are used, tabulated for certain values of  $a$  and  $b$ . The table is also used for any values of  $\mu$  and  $\sigma$  by standardizing the value and using the table (or software).

$$z = \frac{X - \mu}{\sigma}$$

**Notation**

When  $X$  is a normal rv with mean  $\mu$  and variance  $\sigma^2$ ,

$$X \sim N(\mu, \sigma)$$

$$z = \frac{X - \mu}{\sigma}$$

$$P(Z < z) = \Phi(z)$$

## Standard Normal Example

With  $z$ -scores:

- (1)  $P(Z < 1)$
- (2)  $P(Z > 1)$
- (3)  $P(Z < -1)$
- (4)  $P(Z > -1)$
- (5)  $P(-1 < Z < 1)$
- (6)  $z$ -score for top 1%
- (7)  $z$ -score for  $Q1$
- (8)  $z$ -score for  $Q3$

## Normal Example I

Suppose that the diameter at breast height (in.) of trees of a certain type is normally distributed with mean 8.8 and standard deviation 2.8 ( $\mu = 8.8$ ,  $\sigma = 2.8$ ) ( $X \sim N(8.8, 2.8)$ ). Calculate the following:

- (1) Probability the diameter of a randomly selected tree will be at least 10", exceed 10"
- (2) Probability the diameter of a randomly selected tree will exceed 20"
- (3) Probability the diameter of a randomly selected tree will be between 5" and 10"
- (4) Widest 8% are wider than what diameter
- (5) If four trees are selected at random, what is the probability that at least one has a diameter exceeding 10"

## Normal Example II (bahaha)

Spongeboob

## Standard Normal Solutions I

With  $z$ -scores and StatDistributions.com:

- (1)  $P(Z < 1) = \Phi(1) = 0.841345$ : input  $z$ -score, left tail
- (2)  $P(Z > 1) = 1 - \Phi(1) = 0.158655$ : input  $z$ -score, right tail
- (3)  $P(Z < -1) = \Phi(-1) = 0.158655$ : input  $z$ -score, left tail
- (4)  $P(Z > -1) = 1 - \Phi(-1) = 0.841345$ : input  $z$ -score, right tail
- (5)  $P(-1 < Z < 1) = \Phi(1) - \Phi(-1) = 0.682689$ : input  $z$ -score, two tails

## Standard Normal Solutions II

- (6)  $z$  for top 1% is the same as the bottom 99% (or 99<sup>th</sup> percentile).  $z_{0.99} = 2.326348$ : input 0.01 in the p-value, right tail OR 0.99 in p-value and left tail
- (7)  $z$  for  $Q1$  is  $z_{0.25} = -0.67449$ : input 0.25 in the p-value, left tail
- (8)  $z$  for  $Q3$  is  $z_{0.75} = 0.67449$ : input 0.75 in the p-value, left tail

*Empirical Rule derivation:*

(9)  $P(-2 < Z < 2) = 0.9545$ : input  $z$ -score, two tail

(10)  $P(-3 < Z < 3) = 0.9973$ : input  $z$ -score, two tails

### Normal Solutions I part 1

(1)  $P(X < 10) = P\left(Z < \frac{10-8.8}{2.8}\right) = \phi(0.43) = 0.665882$ ,  $P(X > 10) = P\left(Z > \frac{10-8.8}{2.8}\right) = 1 - \phi(0.43) = 0.334118$

(2)  $P(X > 20) = P\left(Z > \frac{20-8.8}{2.8}\right) = 1 - \phi(4) = 0.000032$

(3)  $P(5 < X < 10) = P\left(\frac{10-8.8}{2.8} < Z < \frac{5-8.8}{2.8}\right) = \phi(0.43) - \phi(-1.36) = 0.578515$

### Normal Solutions I part 2

(4) Find  $z$  for top 8% (same as bottom 92%).  $z_{0.92} = 1.405072$ . Use  $z$ -score equation and solve for  $x$ .  
 $z = \frac{X-\mu}{\sigma} \Rightarrow x = z\sigma + \mu$ .  $x = (1.41)(2.8) + 8.8 = 12.748$

(5) Find the probability for one item first:  $P(X > 10) = 0.334118$ . Now, since they are independent, if *four* trees are selected at random, what is the probability that at least one has a diameter exceeding 10".  
 $P(X_4 > 1) = 1 - (X_4 = 0) = 1 - (0.334118)^4 = 0.987538$

### Normal Solutions II

There is no solution for a normal Spongebob... he will be crazy forever... forever... forever...