## ARTICLE

# Proposed Standard Weight $\left(W_{s}\right)$ Equation and Length Categories for Utah Chub 

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#### Abstract

Condition indices, such as relative weight ( $W_{r}$ ), provide a simple method for comparing length-weight relationships among populations. However, no standard weight $\left(W_{s}\right)$ equation has been developed for Utah Chub Gila atraria, a species of important management focus in the Intermountain West. We obtained length-weight data for 30,541 Utah Chub from 24 populations in Idaho, Montana, Utah, and Wyoming. We used the regression line percentile (RLP), linear empirical percentile (EmP), and quadratic EmP methods to develop average (50th percentile) and above average (75th percentile) $W_{s}$ equations. Additionally, Froese's method was used to develop another $W_{s}$ equation for Utah Chub. Length-related biases were detected in $W_{s}$ equations developed using the RLP, 50th percentile quadratic EmP, and Froese methods. The linear EmP $W_{s}$ equations did not exhibit length-related biases for the 50th and 75 th percentiles. We propose using the 75th percentile linear EmP $W_{s}$ equation for Utah Chub between 90 and 410 mm TL. The EmP 75th percentile equation was $\log _{10}\left(W_{s}\right)=-4.938+3.031 \cdot \log _{10}(\mathrm{TL})$, where $W_{s}$ is weight in grams and TL is in millimeters. The English equivalent of this equation is $\log _{10}\left(W_{s}\right)=-3.335+3.031 \cdot \log _{10}(\mathrm{TL})$, where $W_{s}$ is weight in pounds and TL is in inches for 4-16-in Utah Chub. Additionally, we propose that minimum TLs of 100 mm ( $\mathbf{4} \mathbf{~ i n ; ~ s t o c k ) , ~} 200 \mathrm{~mm}$ (8 in; quality), 250 mm ( 10 in ; preferred), 300 mm ( 12 in ; memorable), and 380 mm ( 15 in ; trophy) be used to calculate proportional size distribution (PSD) indices. Better understanding Utah Chub populations using $\boldsymbol{W}_{r}$ and PSDs will aid managers in assessing management strategies (e.g., biological controls) focused on Utah Chub.


Utah Chub Gila atraria is a nongame cyprinid species native to the Snake River and Lake Bonneville basins. Its distribution has expanded to other waters in the western

United States through intentional and unintentional releases (e.g., bait fish introductions; Sigler and Sigler 1987; Rahel 2004). In systems where it is nonnative, it has
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the ability to expand its population quickly once established (Page and Burr 1991; Winters et al. 2017). Typically, Utah Chub are found in lentic systems, where they often compete with salmonids for prey resources (Teuscher and Luecke 1996; Winters and Budy 2015). Utah Chub is a particularly good competitor with salmonids in unnatural or disturbed systems, such as reservoirs (Johnson and Belk 2007). In systems with high Utah Chub densities, salmonid growth, abundance, and survival often decrease (Schneidervin and Hubert 1987; Teuscher and Luecke 1996; Winters and Budy 2015; Winters et al. 2017). A better understanding of the ecology of Utah Chub will aid fishery managers in monitoring trends in Utah Chub populations and in evaluating management actions (e.g., Ward et al. 2008; Roth et al. 2020).

Fisheries assessment tools were originally developed for sport fish populations (Gabelhouse 1984), but the same indices are being applied to nongame species with the goal of providing insight on their population dynamics (Bister et al. 2000; Didenko et al. 2004; Ogle and Winfield 2009). Condition indices provide a simple measure of comparing length-weight relationships among populations and provide insight on the body condition of fish (Brown and Murphy 1991). A commonly used index to assess the condition of fish is relative weight ( $W_{r}$; Wege and Anderson 1978; Neumann et al. 2012). To calculate $W_{r}$, a standard weight $\left(W_{s}\right)$ equation must be developed for the particular species. Once a $W_{s}$ equation is available, $W_{r}$ is calculated as

$$
W_{r}=\left(W / W_{s}\right) \times 100
$$

where $W$ is the weight of the fish and $W_{s}$ is the standard weight for a fish of the same length (Wege and Anderson 1978). Standard weight equations are estimated using various techniques, particularly the regression line percentile (RLP) and empirical percentile (EmP) methods (Murphy et al. 1990; Gerow et al. 2005). The RLP has been the most widely used method for developing $W_{s}$ equations and uses the 75th percentile of the mean weights estimated among populations as the basis for the $W_{s}$ equation (Murphy et al. 1990). However, Gerow et al. (2005) found that length-related biases can occur with the development of $W_{s}$ equations using the RLP method and suggested that the EmP method may be a more appropriate alternative. The EmP technique is based on the 75th percentile of the observed weights by $10-\mathrm{mm}$ length increments as opposed to the RLP technique, which uses weights estimated from regression models. Selection of the most appropriate method (i.e., RLP versus EmP) for development of the $W_{s}$ equation has not entirely been resolved (Gerow 2010; Ranney et al. 2010). Another option for developing a $W_{s}$ equation is the method proposed by Froese (2006). Froese's method uses the geometric mean and mean produced
from all available weight-length estimates to represent the mean weight for the focal species.

The objective of this study was to develop a $W_{s}$ equation for Utah Chub based on the RLP, EmP, and Froese methods and to establish minimum TLs that can be used to estimate length structure following the five-cell system described by Gabelhouse (1984). This information will provide managers with the necessary tools to effectively monitor Utah Chub populations.

## METHODS

Data summarization and standard weight equation.-Length-weight data for Utah Chub were acquired from state, federal, and university biologists across the known geographic distribution of Utah Chub. Personnel providing data were asked to include the TL, weight, sex, date sampled, and sampling gear used. Data sets were screened and evaluated prior to analysis. First, data collected during June were excluded to minimize length-weight discrepancies associated with the development of mature gonads and spawning (Graham 1961; Ogle and Winfield 2009). Next, outliers were removed by plotting Utah Chub $\log _{10}(\mathrm{TL})$ against $\log _{10}$ (weight) across years for each water body (Bister et al. 2000; Ogle and Winfield 2009). Individual fish that exceeded the $99 \%$ prediction interval from each water body were considered outliers and removed from the data set (Bister et al. 2000; Ogle and Winfield 2009). Utah Chub that were one of a few individuals (i.e., three or less) that were substantially shorter or longer than the majority of individuals in the data set were also considered outliers and removed from the analysis (Ogle and Winfield 2009). This resulted in $0.25 \%$ of the total number of individuals being removed and had no effect on the number or composition of length categories used in subsequent analyses. After removing outliers, water bodies with less than 20 fish were removed from the data set. Furthermore, water bodies with a log-transformed length-weight regression with a coefficient of determination less than 0.90 or with a slope (b) beyond $2.5-3.5$ were removed from the analysis (Ogle and Winfield 2009). We further analyzed the estimated intercepts $\left(\log _{10}[a]\right)$ regressed against the estimated slopes for each water body to evaluate any divergent patterns among populations. No additional data were excluded from analysis.

The data were then randomly separated into "developmental" and "validation" data sets. The developmental data set was used to compute the $W_{s}$ equation, whereas the validation data set was used to evaluate the proposed $W_{s}$ equation and examine any length-related bias (Gerow et al. 2005; Ogle and Winfield 2009). The developmental data set contained 18,130 individuals representing 24 populations. Populations assigned to the developmental data set had 1-9 years of data and varied from 24 to 7,709
individuals (Table 1). Populations assigned to the validation data set had large samples ( $\geq 500$ fish $)$ and at least 2 years of data. These criteria were used to identify populations that were large enough to be further partitioned into the developmental or validation data sets. Nine of the 24 populations met these criteria. Years within populations were randomly assigned to either the developmental or validation data set. Consecutive years placed in a data set were avoided while still assigning a large proportion to the developmental data set (Ogle and Winfield 2009). For example, Blackfoot Reservoir, Idaho, contained 3 years of data and 500 fish. Therefore, Blackfoot Reservoir met the requirements for inclusion in the validation data set. Three-hundred eighty-three Utah Chub were sampled in 2009 and 2012 from Blackfoot Reservoir; years were combined and included in the developmental data set. The remaining year was assigned to the validation data set (Table 1). At the end of this process, the validation data set contained 12,280 individuals representing nine populations. Populations assigned to the validation data set had $1-8$ years of data and varied from 117 to 5,733 individuals (Table 1). Hereafter, "population" refers to the combined population-years of a water body. For example, data from 2009 and 2012 in Blackfoot Reservoir were combined to represent a single "population" in the developmental data set. Differences in mean slope and mean intercept values between the developmental and validation data sets were tested using a two-sample $t$-test (Ogle and Winfield 2009). The number of populations required for a robust $W_{s}$ equation from the developmental data set was estimated using a bootstrap technique (Brown and Murphy 1996). Specifically, slopes from $\log _{10}(\mathrm{TL})$ against $\log _{10}$ (weight) regressions (by population) were used as modeling parameters and were randomly selected with replacement for 300 iterations. Sample variance of the slope was calculated for each incremental sample size ( $n=2-24$ populations). A sample variance of less than 0.002 was selected as the decision criterion for the minimum number of populations required for the $W_{s}$ equation (Brown and Murphy 1996).

Following the compilation of the data sets, an estimate of a suitable length distribution was developed prior to the derivation of $W_{s}$ equations, given that measurements from small fish tend to have low precision and accuracy (Rypel and Richter 2008; Neumann et al. 2012). The minimum and maximum TLs were established using the developmental data set. Minimum TL was established for the RLP method by plotting the variance-to-mean ratio for $\log _{10}$ (weight) by $10-\mathrm{mm}$ intervals and described when the ratio fell below 0.01 (Murphy et al. 1990; Figure 1). Maximum TL was established for the EmP method when at least three populations contained the largest size-class (Gerow et al. 2005; Table 2).

Standard weight equations were estimated using the RLP, EmP, and Froese methods (Murphy et al. 1990;

Gerow et al. 2005; Froese 2006). For RLP and EmP methods, the 50th ( $W_{s 50}$ ) and 75th ( $W_{s 75}$ ) percentiles were estimated using the mean of the predicted weights in each $10-\mathrm{mm}$ length interval from individual populations produced from the weight-length relationships (Brown and Murphy 1996). Using quadratic regression, the third quartile of mean weight by TL was used to develop the $W_{s}$ equation (Gerow et al. 2005). Froese's method ( $W_{\text {smean }}$ ) was assessed by calculating mean intercept and mean slope from length-weight regressions for each population in the developmental data set (Froese 2006).

Length-related biases of the $W_{s}$ equations were assessed using the residual analysis, Willis method (Willis et al. 1991), and empirical quartiles (EmpQ) method (Gerow et al. 2004). For the residual analysis, residuals of the $W_{s}$ equation were visually examined for evident patterns. The Willis method used a chi-square test to identify whether the proportion of populations with a significant positive slope (i.e., TL versus $W_{r}$ using the validation data) was equal to the proportion with a significant negative slope. To assess the length-related bias of the quadratic EmP equations, the EmpQ method used the validation data set to evaluate whether the slope of the third quartile of the mean weight (standardized by $W_{s}$ midpoint length intervals of 10 mm ) was zero. We used the Fisheries Stock Assessment package (version 0.0-14; Ogle 2009) using R software (version 4.0.3; R Core Team 2020) for all analyses.

Length categories for stock size distribution indices.We used the criteria developed by Gabelhouse (1984) to establish a minimum TL of Utah Chub for each of the five-cell length categories (i.e., stock, quality, preferred, memorable, and trophy). For development of the minimum TL categories, we excluded all Utah Chub from Jackson Lake, Wyoming. The longest specimen from Jackson Lake was over 115 mm larger than the next largest specimen from any other population, and Jackson Lake contained the 12 largest individuals in the data set. Using the largest specimen from Jackson Lake would have resulted in minimum TLs that are too large for most Utah Chub populations.

## RESULTS AND DISCUSSION

Data were obtained from 24 Utah Chub populations that spanned over 40 years (1977-2017) from four statesIdaho, Montana, Utah, and Wyoming. All populations were from lentic systems except for a population in Billingsley Creek, Idaho. The entire data set included 30,410 Utah Chub that varied in length from 65 to 612 mm (mean $\pm 95 \% \mathrm{CI}=232 \pm 0.7 \mathrm{~mm}$ ) and from 3 to $2,168 \mathrm{~g}$ (mean $\pm 95 \% \mathrm{CI}=202 \pm 1.9 \mathrm{~g}$; Table 1). For all populations used in the developmental and validation data sets, mean slope was 3.127 ( $95 \% \mathrm{CI}=3.072-3.182$ ). When

TABLE 1. Populations used to develop standard weight equations for Utah Chub in Idaho (ID), Montana (MT), Utah (UT), and Wyoming (WY) by gear type used ( $\mathrm{EF}=$ electrofishing; $\mathrm{GN}=$ gill net; $\mathrm{MNT}=$ minnow trap; $\mathrm{TN}=$ trap net; UNK $=$ unknown sampling gear $)$, data set use $(\mathrm{D}=$ developmental; $\mathrm{V}=$ validation), years, sample size $(n)$, minimum ( $\min$ ) and maximum (max) TL and weight, and estimated intercept ( $\log _{10}[a]$ ) and slope (b) of the $\log _{10}$ transformed length-weight data.

| State | Population | Use | Gear | Years | $n$ | TL (mm) |  | Weight (g) |  | Length-weight equation |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Min | Max | Min | Max | $\log _{10}(a)$ | $b$ | $R^{2}$ |
| ID | American Falls | D | GN | 1997 | 129 | 137 | 406 | 30 | 1,000 | -4.706 | 2.950 | 0.989 |
|  | Reservoir |  |  |  |  |  |  |  |  |  |  |  |
|  | Billingsley Creek | D | EF | 2006 | 60 | 65 | 285 | 3 | 277 | -4.928 | 3.012 | 0.995 |
|  | Blackfoot Reservoir | D | $\begin{gathered} \mathrm{GN}, \\ \mathrm{TN} \end{gathered}$ | 2009, 2012 | 383 | 135 | 495 | 32 | 1,780 | -5.029 | 3.078 | 0.993 |
|  |  | V | GN | 2015 | 117 | 120 | 435 | 28 | 1,118 | -4.561 | 2.886 | 0.991 |
|  | Henrys <br> Lake | D | GN | $\begin{aligned} & \text { 2005, 2007, 2010, 2014, } \\ & 2015,2017 \end{aligned}$ | 7,709 | 120 | 373 | 18 | 758 | -4.998 | 3.044 | 0.981 |
|  |  | V | GN | $\begin{aligned} & \text { 2004, 2006, 2008, 2009, } \\ & 2012,2013,2016 \end{aligned}$ | 5,733 | 92 | 370 | 9 | 654 | -5.099 | 3.089 | 0.983 |
|  | Island Park Reservoir | D | GN | 2008, 2016 | 146 | 82 | 360 | 5 | 626 | -5.398 | 3.206 | 0.991 |
|  | Ririe | D | GN | 2013, 2016 | 299 | 150 | 422 | 37 | 890 | -4.823 | 2.971 | 0.963 |
|  | Reservoir | V | UNK | 2012, 2014 | 240 | 145 | 390 | 40 | 635 | -5.045 | 3.063 | 0.912 |
| $\begin{aligned} & \text { ID/ } \\ & \text { WY } \end{aligned}$ | Palisades Reservoir | D | UNK | 2012 | 189 | 150 | 391 | 36 | 885 | -5.698 | 3.329 | 0.974 |
| MT | Hebgen Lake | D | UNK | $\begin{aligned} & \text { 1998, 2006, 2007, 2012, } \\ & 2015 \end{aligned}$ | 1,857 | 127 | 404 | 23 | 835 | -4.976 | 3.037 | 0.971 |
|  |  | V | UNK | 1999, 2002, 2009, 2011 | 1,939 | 145 | 406 | 27 | 816 | -5.202 | 3.127 | 0.972 |
| UT | Fish Lake | D | GN | 2014, 2015 | 85 | 107 | 285 | 15 | 319 | -5.145 | 3.111 | 0.969 |
|  | Lost Creek | D | GN | 2012, 2014 | 842 | 132 | 283 | 26 | 303 | -4.996 | 3.062 | 0.94 |
|  | Reservoir | V | GN | 2011, 2013 | 626 | 133 | 265 | 29 | 257 | -5.999 | 3.491 | 0.90 |
|  | Minersville Reservoir | D | GN | 2015 | 30 | 162 | 278 | 41 | 305 | -5.806 | 3.383 | 0.91 |
|  | Otter Creek Reservoir | D | GN | 2015 | 24 | 140 | 340 | 42 | 511 | -4.688 | 2.934 | 0.99 |
|  | Piute <br> Reservoir | D | GN | 2010, 2011, 2014 | 31 | 138 | 283 | 33 | 312 | -5.203 | 3.110 | 0.97 |
|  | Rockport Reservoir | D | GN | 2012, 2015 | 70 | 158 | 334 | 45 | 599 | -5.396 | 3.241 | 0.96 |
|  | Scofield Reservoir | D | $\begin{aligned} & \text { GN, } \\ & \text { MNT } \end{aligned}$ | 2012 | 1,084 | 94 | 336 | 8 | 543 | -5.404 | 3.219 | 0.98 |
|  |  | V | $\begin{aligned} & \mathrm{GN}, \\ & \mathrm{MNT} \end{aligned}$ | 2011 | 745 | 94 | 325 | 9 | 469 | -5.436 | 3.233 | 0.99 |
|  | Strawberry Reservoir | D | UNK | $\begin{aligned} & 2002,2006,2008,2009 \\ & 2012,2013,2016 \end{aligned}$ | 2,216 | 92 | 435 | 9 | 1,150 | -5.495 | 3.269 | 0.99 |
|  |  | V | UNK | $\begin{aligned} & 2000,2001,2005,2008 \\ & 2009,2013,2014 \end{aligned}$ | 1,666 | 95 | 361 | 9 | 860 | -5.532 | 3.280 | 0.99 |
|  | Woodruff Reservoir | D | UNK | 1997 | 356 | 150 | 328 | 36 | 540 | -5.789 | 3.385 | 0.99 |
| $\begin{gathered} \text { UT/ } \\ \text { WY } \end{gathered}$ | Flaming Gorge Reservoir | D | UNK | 2001, 2003, 2009 | 208 | 191 | 361 | 86 | 662 | -5.527 | 3.265 | 0.95 |

TABLE 1. Continued.

| State | Population | Use | Gear | Years | $n$ | TL (mm) |  | Weight (g) |  | Length-weight equation |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Min | Max | Min | Max | $\log _{10}(a)$ | $b$ | $R^{2}$ |
| WY | Fontenelle | D | UNK | 1998, 1991, 1993, 2004 | 271 | 107 | 285 | 14 | 771 | -5.085 | 3.085 | 0.98 |
|  | Reservoir | V | UNK | 1984, 1992, 1997, 2009 | 314 | 104 | 366 | 9 | 671 | -5.515 | 3.261 | 0.99 |
|  | Jackson Lake | D | UNK | $\begin{aligned} & 1978,1990,1993,1995- \\ & 1997,2000,2002,2004 \\ & 2006,2007,2009 \end{aligned}$ | 1,273 | 127 | 612 | 27 | 2,168 | -4.992 | 3.024 | 0.99 |
|  |  | V | UNK | $\begin{array}{r} 1977,1979,1980,1998, \\ 1999,2003,2005,2008 \end{array}$ | 900 | 132 | 475 | 32 | 1,139 | -4.902 | 2.989 | 0.99 |
|  | Lower Slide Lake | D | UNK | $\begin{aligned} & 1998-2000,2002-2007, \\ & 2009-2012 \end{aligned}$ | 302 | 135 | 422 | 27 | 776 | -4.566 | 2.838 | 0.92 |
|  | Naughton <br> Plant <br> Pond | D | UNK | 2009 | 107 | 124 | 323 | 23 | 376 | -4.860 | 2.944 | 0.98 |
|  | Two Ocean Lake | D | UNK | 1981, 1995, 2009 | 153 | 135 | 376 | 32 | 549 | -5.236 | 3.110 | 0.99 |
|  | Viva <br> Naughton Reservoir | D | UNK | 1981, 1983, 1989, 1993 | 306 | 137 | 328 | 27 | 553 | -5.290 | 3.173 | 0.96 |



FIGURE 1. Variance-to-mean ratio of $\log _{10}$ (weight) by TL at $10-\mathrm{mm}$ intervals for Utah Chub used in the developmental data set. Dashed line indicates the value of 0.01 .
$b$ was regressed on $\log _{10}(a)$ for the populations in the developmental data set $\left(\log _{10}[a]=-2.319[b]+2.058 ; r^{2}=\right.$ $0.98)$ and populations in the validation data set $\left(\log _{10}[a]=\right.$ $\left.-2.323[b]+2.081 ; r^{2}=0.99\right)$, no obvious outliers were
observed. Also, differences in mean slope $(t=0.684, \mathrm{df}=$ $31, P=0.50$ ) and mean intercept ( $t=-0.607, \mathrm{df}=31, P=$ 0.55 ) between the developmental and validation data sets were not significant. Proportions between the

TABLE 2. Number of populations and number of individual Utah Chub in the developmental data set used to develop the standard weight ( $W_{s}$ ) equation per $10-\mathrm{mm}$ length-class (TL). Length-classes marked with an asterisk were not used in estimating the $W_{s}$ equation of the linear empirical percentile due to low sample size $(n<3)$.

| Length-class (mm) | Populations | Individuals | Length-class (mm) | Populations | Individuals |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 70* | 1 | 2 | 330 | 17 | 359 |
| 80* | 2 | 3 | 340 | 14 | 225 |
| 90 | 3 | 3 | 350 | 12 | 206 |
| 100 | 3 | 36 | 360 | 12 | 206 |
| 110 | 6 | 52 | 370 | 10 | 167 |
| 120 | 5 | 31 | 380 | 7 | 140 |
| 130 | 10 | 70 | 390 | 5 | 94 |
| 140 | 15 | 545 | 400 | 3 | 42 |
| 150 | 20 | 1,331 | 410 | 3 | 43 |
| 160 | 22 | 1,340 | 420 | 4 | 17 |
| 170 | 20 | 1,088 | 430* | 1 | 10 |
| 180 | 17 | 715 | 440 | 3 | 7 |
| 190 | 22 | 911 | 450* | 1 | 3 |
| 200 | 23 | 1,120 | 460* | 1 | 7 |
| 210 | 22 | 1,274 | 470* | 1 | 4 |
| 220 | 21 | 1,046 | 480* | 1 | 1 |
| 230 | 22 | 869 | 490* | 1 | 4 |
| 240 | 24 | 763 | 500* | 2 | 2 |
| 250 | 23 | 718 | 510* | 1 | 2 |
| 260 | 24 | 640 | $520 *$ | 1 | 2 |
| 270 | 21 | 640 | 530* | 1 | 1 |
| 280 | 23 | 734 | 540* | 1 | 1 |
| 290 | 18 | 691 | 550* | 1 | 2 |
| 300 | 16 | 756 | 560* | 1 | 1 |
| 310 | 17 | 728 | 570* | 1 | 1 |
| 320 | 19 | 476 | 610* | 1 | 1 |

TABLE 3. Parameter estimates for Utah Chub standard weight ( $W_{s}$ ) equation using regression line percentile (RLP), linear empirical percentile (EmP), quadratic EmP, and Froese methods ( $W_{s 50}=50$ th percentile; $W_{s 75}=75$ th percentile; EmpQ $=$ empirical quartiles).

| Equation | Parameters |  |  | Willis test |  |  | EmpQ test |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\log _{10}(a)$ | $b_{\text {linear }}$ | $b_{\text {quadratic }}$ | Negative | Positive | $P$ | $P_{\text {linear }}$ | $P_{\text {quadratic }}$ |
| RLP $W_{s 50}$ | -5.1454 | 3.1220 | NA | 3 | 4 | 1.0000 | 0.0078 | 0.0015 |
| $\mathrm{EmP} W_{s 50}$ | -6.4431 | 4.2338 | -0.2442 | 3 | 5 | 0.7266 | 0.2682 | 0.0471 |
| EmP $W_{s 50}$ | -5.1337 | 3.0987 | NA | 3 | 4 | 1.0000 | 0.0773 | 0.0008 |
| RLP $W_{s 75}$ | -5.1454 | 3.1220 | NA | 3 | 4 | 1.0000 | 0.0373 | 0.5710 |
| EmP $W_{s 75}$ | -4.4103 | 2.5728 | 0.0985 | 2 | 6 | 0.2891 | 0.4697 | 0.0858 |
| EmP $W_{s 75}$ | -4.9386 | 3.0308 | NA | 2 | 6 | 0.2891 | 0.3561 | 0.2217 |
| Froese $W_{\text {smean }}$ | -5.1681 | 3.1158 | NA | 3 | 4 | 1.0000 | 0.0022 | 0.0005 |

developmental and validation data sets with a slope statistically less than, equal to, or greater than 3.0 were not significantly different $\left(\chi^{2}=0.060, \mathrm{df}=2, P=0.97\right)$. Based on the bootstrap analysis, variability among slopes was low and the minimum number of populations necessary for the developmental data set (sample variance $<0.002$ ) was
eight populations. The minimum TL for the $W_{s}$ equation was 90 mm based on the variance-to-mean ratio for RLP (Figure 1) and 90 mm for EmP due to the low sample size ( $n<3$ populations; Table 2). Due to the low number of populations with fish in the $450-\mathrm{mm}$ interval ( $n<$ 3), the maximum TL for the $W_{s}$ equation was 440 mm for


FIGURE 2. Distribution of standardized residuals of the regression plotted against TL category midpoints of Utah Chub obtained by model fit of the 50th and 75th percentiles for the linear empirical percentile (EmP), quadratic empirical percentile (EmP-Q), and regression line percentile (RLP) methods.

EmP. The maximum TL for the $W_{s}$ equation was 440 mm for the RLP to match that used in the EmP. Although the developmental data set included fish from 90 to 440 mm , the $W_{s}$ equation is only applicable to $90-410-\mathrm{mm}$ fish due to the availability of size-classes present in the validation data set.

The RLP $W_{s 50}$ equation exhibited significant lengthrelated bias when evaluated with the EmpQ method (Table 3 ) and overpredicted the $W_{r}$ across all lengths. Residuals of
the RLP $W_{s 50}$ equation also exhibited nonlinearity (Figure 2). Residuals from the linear EmP $W_{s 50}$ equation exhibited no obvious pattern, and a slight length-related bias was detected with the EmpQ method. The quadratic EmP $W_{s 50}$ exhibited a length-related bias (Table 3). Residuals of the RLP $W_{s 75}$ equation exhibited nonlinearity (Figure 2), and length-related bias was detected with the EmpQ method (Table 3). The RLP $W_{s 75}$ equation tended to overpredict $W_{r}$ with length (Figure 3). The EmP $W_{s 75}$ equation showed no


FIGURE 3. Standardized mean Utah Chub weights compared to TL category midpoints with the weighted quadratic fit for the Froese method and 50th and 75 th percentiles for the linear empirical percentile ( EmP ), quadratic empirical percentile (EmP-Q), and regression line percentile (RLP) methods for establishing standard weight equations.
length-related bias with the Willis or EmpQ method for the linear and quadratic terms. Froese $W_{\text {smean }}$ exhibited significant length-related bias with the EmpQ method (Table 2) and overpredicted the $W_{r}$ with increasing length (Figure 3).

The longest fish observed in the data set excluding Jackson Lake was a $495-\mathrm{mm}$ individual from Blackfoot Reservoir, Idaho. Using this length, the proposed minimum TLs are 100 mm ( 4 in ; stock), 200 mm ( 8 in ; quality), $250 \mathrm{~mm}(10 \mathrm{in}$; preferred), 300 mm (12 in; memorable), and 380 mm ( 15 in ; trophy).

Length-weight data are rarely collected for nongame species, and Murphy et al. (1990) suggested 50 populations for the development of a $W_{s}$ equation. We did not collect data from the suggested 50 populations but were successful in obtaining data across the geographical distribution of the species. Variance among populations was low, suggesting that 50 populations were not needed. The linear EmP equation(s) provided the best fit to the data and did not result in length-related bias. We propose using the 75 th percentile linear EmP $W_{s}$ equation for $100-410-\mathrm{mm}$ Utah Chub: $\log _{10}\left(W_{s}\right)=-4.938+3.031 \cdot \log _{10}(\mathrm{TL})$, where $W_{s}$ is weight in grams and TL is in millimeters. The English equivalent of this equation is $\log _{10}\left(W_{s}\right)=-3.335+3.031 \cdot \log _{10}(\mathrm{TL})$, where $W_{s}$ is weight in pounds and TL is in inches for 4-16in Utah Chub. The linear EmP $W_{s 75}$ equation showed no length-related bias and can be a valuable tool for monitoring Utah Chub populations.

Despite the potential negative influence Utah Chub may have on sport fish populations (Teuscher and Luecke 1996; Winters and Budy 2015; Roth et al. 2020), limited information exists regarding their life history and population dynamics. Relative weights coupled with other population metrics (e.g., age structure, growth; Griffin et al. 2017; Roth et al. 2020) of Utah Chub could increase basic knowledge on their population dynamics and ultimately improve sport fish management. The use of the proposed length categories in conjunction with $W_{r}$ could be useful for providing additional information on factors influencing Utah Chub populations. Relative weights could be used as a simple, less invasive tool to monitor changes in Utah Chub populations (e.g., compared to growth evaluations) and evaluate management strategies across the geographical range of Utah Chub. Temporal trends of $W_{r}$ for a specific population may also provide understanding of the abiotic (e.g., temperature, oxygen) and biotic (e.g., intraspecific and interspecific competition, predation) factors potentially regulating their populations. The $W_{r}$ values estimated for this nongame species provide an additional tool for management efforts aimed at enhancing or reducing densities.

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